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Using the Mean Absolute Percentage Error for Regression Models

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Abstract. We study in this paper the consequences of using the Mean Absolute Percentage Error (MAPE) as a measure of quality for regression models. We show that finding the best model under the MAPE is equivalent to doing weighted Mean Absolute Error (MAE) regression. We show that universal consistency of Empirical Risk Minimization remains possible using the MAPE instead of the MAE.

1 Introduction

We study in this paper the classical regression setting in which we assume given a random pair $Z = (X, Y)$ with values in $\mathcal{X} \times \mathbb{R}$, where \mathcal{X} is a metric space. The goal is to learn a mapping g from \mathcal{X} to \mathbb{R} such that $g(X) \simeq Y$. To judge the quality of the regression model g , we need a quality measure. While the traditional measure is the quadratic error, in some applications, a more useful measure of the quality of the predictions made by a regression model is given by the mean absolute percentage error (MAPE). For a target y and a prediction p , the MAPE is

$$l_{MAPE}(p, y) = \frac{|p - y|}{|y|},$$

with the conventions that for all $a \neq 0$, $\frac{a}{0} = \infty$ and that $\frac{0}{0} = 1$. The MAPE-risk of g is then $L_{MAPE}(g) = \mathbb{E}(l_{MAPE}(g(X), Y))$.

We are interested in the consequences of choosing the best regression model according to the MAPE as opposed to the Mean Absolute Error (MAE) or the Mean Square Error (MSE), both on a practical point of view and on a theoretical one. On a practical point of view, it seems obvious that if g is chosen so as to minimize $L_{MAPE}(g)$ it will perform better according to the MAPE than a model selected in order to minimize L_{MSE} (and worse according to the MSE). The practical issue is rather to determine how to perform this optimization: this is studied in Section 3. On a theoretical point of view, it is well known (see e.g. [4]) that consistent learning schemes can be obtained by adapting the complexity of the model class to the data size. As the complexity of a class of models is partially dependent on the loss function, using the MAPE instead of e.g. the MSE has some implications that are investigated in this paper, in Section 4. The following Section introduces the material common to both parts of the analysis.

2 General setting

We use a classical statistical learning setting as in e.g. [4]. We assume given N independently distributed copies of Z , the training set, $D = (Z_i)_{1 \leq i \leq N} = (X_i, Y_i)_{1 \leq i \leq N}$. Given a loss function l from \mathbb{R}^2 to $\mathbb{R}^+ \cup \{\infty\}$, we define the risk of a predictor g , a (measurable) function from \mathcal{X} to \mathbb{R} as the expected loss, that is $L_l(g) = \mathbb{E}(l(g(X), Y))$. The empirical risk is the empirical mean of the loss computed on the training set, that is:

$$\hat{L}_l(g)_N = \frac{1}{N} \sum_{i=1}^N l(g(X_i), Y_i). \quad (1)$$

In addition to l_{MAPE} defined in the Introduction, we use $l_{MAE}(p, y) = |p - y|$ and $l_{MSE}(p, y) = (p - y)^2$.

3 Practical issues

3.1 Optimization

On a practical point of view, the problem is to minimize $\hat{L}_{MAPE}(g)_N$ over a class of models G_N , that is to solve¹

$$\hat{g}_{MAPE, N} = \arg \min_{g \in G_N} \frac{1}{N} \sum_{i=1}^N \frac{|g(X_i) - Y_i|}{|Y_i|}.$$

Optimization wise, this is simply a particular case of *median regression* (which is in turn a particular case of *quantile regression*). Indeed, the quotient by $\frac{1}{|Y_i|}$ can be seen as a fixed weight and therefore, any quantile regression implementation that supports instance weights can be use to find the optimal model². Notice that when G_N corresponds to linear models, the optimization problem is a simple *linear programming* problem that can be solved by e.g. interior point methods [2].

3.2 An example of typical results

We verified on a toy example (the car data set from [3]) the effects of optimizing the MAPE, the MAE and the MSE for a simple linear model: the goal is to predict the distance taken to stop from the speed of the car just before breaking. There are only 50 observations, the goal being here to illustrate the effects of changing the loss function. The results on the training set³ are summarized in Table 1. As expected, optimizing for a particular loss function leads to the best

¹We are considering here the empirical risk minimization, but we could of course include a regularization term. That would not modify the key point which is the use of the MAPE.

²This is the case of **quantreg** R package [5], among others.

³Notice that the goal here is to verify the effects of optimizing with respect to different types of loss function, not to claim that one loss function is better than another, something that would be meaningless. We report therefore the empirical risk, knowing that it is an underestimation of the real risk for all loss functions.

empirical model as measured via the same risk (or a related one). In practice this allowed one of us to win a recent `datascience.net` challenge about electricity consumption prediction⁴ which was using the MAPE as the evaluation metric.

Loss function	RMSE	NMAE	MAPE
MSE	0.585	0.322	0.384
MAE	0.601	0.313	0.330
MAPE	0.700	0.343	0.303

Table 1: Empirical risks of the best linear models obtained with the three loss functions. In order to ease the comparisons between the values, we report the Normalized Root MSE, that is the square root of the MSE divided by the standard deviation of the target variable, as well as the Normalized MAE, that is the MAE divided by the median of the target variable.

4 Theoretical issues

On a theoretical point of view, we are interested in the consistency of standard learning strategies when the loss function is the MAPE. More precisely, for a loss function l , we define $L_l^* = \inf_g L_l(g)$, where the infimum is taken over all measurable functions from \mathcal{X} to \mathbb{R} . We also denote $L_{l,G}^* = \inf_{g \in G} L_l(g)$ where G is a class of models. Then a learning algorithm, that is a function which maps the training set $D = (X_i, Y_i)_{1 \leq i \leq N}$ to a model \hat{g}_N , is strongly consistent if $L_l(\hat{g}_N)$ converges almost surely to L_l^* . We are interested specifically by the Empirical Risk Minimization (ERM) algorithm, that is by $\hat{g}_{l,N} = \arg \min_{g \in G_N} \hat{L}_l(g)_N$. The class of models to depend on the data size as this is mandatory to reach consistency.

It is well known (see e.g. [4] chapter 9) that ERM consistency is related to uniform laws of large numbers (ULLN). In particular, we need to control quantities of the following form

$$P \left\{ \sup_{g \in G_N} \left| \hat{L}_{\text{mape}}(g)_N - L_{\text{mape}}(g) \right| > \epsilon \right\}. \quad (2)$$

This can be done via covering numbers or via the Vapnik-Chervonenkis dimension (VC-dim) of certain classes of functions derived from G_N . One might think that general results about arbitrary loss functions can be used to handle the case of the MAPE. This is not the case as those results generally assume a uniform Lipschitz property of l (see Lemma 17.6 in [1], for instance) that is not fulfilled by the MAPE.

4.1 Classes of functions

Given a class of models, G_N , and a loss function l , we introduce derived classes $H(G_N, l)$ given by

$$H(G_N, l) = \{h : \mathcal{X} \times \mathbb{R} \rightarrow \mathbb{R}^+, h(x, y) = l(g(x), y) \mid g \in G_N\},$$

⁴<https://datascience.net/fr/challenge/16/details>

and $H^+(G_N, l)$ given by

$$H^+(G_N, l) = \{h : \mathcal{X} \times \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}^+, h(x, y, t) = \mathbb{I}_{t \leq l(g(x), y)} \mid g \in G_N\}.$$

When this is obvious from the context, we abbreviate the notations into e.g. $H_{N, MAPE}$ for $l = l_{MAPE}$ and for the G_N under study.

4.2 Covering numbers

4.2.1 Supremum covering numbers

Let $\epsilon > 0$, a size p supremum ϵ -cover of a class of positive functions F from an arbitrary set \mathcal{Z} to \mathbb{R}^+ is a finite collection f_1, \dots, f_p of F such that for all $f \in F$

$$\min_{1 \leq i \leq p} \sup_{z \in \mathcal{Z}} |f(z) - f_i(z)| < \epsilon.$$

Then the supremum ϵ -covering number of F , $\mathcal{N}_\infty(\epsilon, F)$, is the size of the smallest supremum ϵ -cover of F . If such a cover does not exist, the covering number is ∞ . While controlling supremum covering numbers of $H(G_N, l)$ leads easily to consistency via a uniform law of large numbers (see e.g. Lemma 9.1 in [4]), they cannot be used with the MAPE without additional assumptions. Indeed, let h_1 and h_2 be two functions from $H_{N, MAPE}$, generated by g_1 and g_2 in G_N . Then

$$\|h_1 - h_2\|_\infty = \sup_{(x, y) \in \mathcal{X} \times \mathbb{R}} \frac{||g_1(x) - y| - |g_2(x) - y||}{|y|}.$$

In general, this quantity will be unbounded as we cannot control the behavior of $g(x)$ around $y = 0$ (indeed, in the supremum, x and y are independent and thus unless G_N is very restricted there is always x and g_1 and g_2 such that $g_1(x) \neq g_2(x) \neq 0$). Thus we have to assume that there is $\lambda > 0$ such that $|Y| \geq \lambda$. This is not needed when using more traditional loss functions such as the MSE or the MAE. Then we have

$$\mathcal{N}_\infty(\epsilon, H(G_N, l_{MAPE})) \leq \mathcal{N}_\infty(\lambda\epsilon, H(G_N, l_{MAE})).$$

4.2.2 L_p covering numbers

L_p covering numbers are similar to supremum covering numbers but are based on a different metric on the class of functions F and are data dependent. Given a data set D , we define

$$\|f_1 - f_2\|_{p, D} = \left(\frac{1}{N} \sum_{i=1}^N |f_1(Z_i) - f_2(Z_i)|^p \right)^{\frac{1}{p}},$$

and derive from this the associated notion of ϵ -cover and of covering number. It's then easy to show that

$$\mathcal{N}_p(\epsilon, H(G_N, l_{MAPE}), D) \leq \mathcal{N}_p(\epsilon \min_{1 \leq i \leq N} |Y_i|, H(G_N, l_{MAE}), D).$$

4.3 Uniform law of large numbers

In order to get a ULLN from a covering number of a class F , one needs a uniform bound on F . For instance, Theorem 9.1 from [4] assumes that there is a value B_F such that for all $f \in F$ and all $z \in \mathcal{Z}$, $f(z) \in [0, B_F]$. With classical loss functions such as MAE and MSE, this is achieved via upper bounding assumptions on both G_N and on $|Y|$. In the MAPE case, the bound on G_N is needed but the upper bound on $|Y|$ is replaced by the lower bound already needed. Let us assume indeed that for all $g \in G_N$, $\|g\|_\infty \leq B_{G_N}$. Then if $|Y| \leq B_Y$, we have $B_{H(G_N, l_{MAE})} = B_{G_N} + B_Y := B_{N,MAE}$, while if $|Y| \geq \lambda$, we have $B_{H(G_N, l_{MAPE})} = 1 + \frac{B_{G_N}}{\lambda} := B_{N,MAPE}$.

Theorem 9.1 from [4] gives then (with $B_{N,l} = B_{H(G_N, l)}$)

$$P \left\{ \sup_{g \in G_N} \left| \widehat{L}_l(g)_N - L_l(g) \right| > \epsilon \right\} \leq 8\mathbb{E} \left(\mathcal{N}_p \left(\frac{\epsilon}{8}, H(G_N, l), D \right) \right) e^{-\frac{N\epsilon^2}{128B_{N,l}^2}}. \quad (3)$$

The expressions of the two bounds above show that B_Y and λ play similar roles on the exponential decrease of the right hand side bound. Loosening the condition on Y (i.e., taking a large B_Y or a small λ) slows down the exponential decrease.

It might seem from the results on the covering numbers that the MAPE suffers more from the bound needed on Y than e.g. the MAE. This is not the case as bounds hypothesis on F are also needed to get finite covering numbers (see the following section for an example). Then we can consider that the lower bound on $|Y|$ plays an equivalent role for the MAPE to the one played by the upper bound on $|Y|$ for the MAE/MSE.

4.4 VC-dimension

A convenient way to bound covering numbers is to use VC-dimension. Interestingly replacing the MAE by the MAPE cannot increase the VC-dim of the relevant class of functions.

Let us indeed consider a set of k points shattered by $H^+(G_N, l_{MAPE})$, (v_1, \dots, v_k) , $v_j = (x_j, y_j, t_j)$. Then for each $\theta \in \{0, 1\}^k$, there is $h_\theta \in H_{N,MAPE}$ such that $\forall j, \mathbb{I}_{t \leq h_\theta(x,y)}(x_j, y_j, t_j) = \theta_j$. Each h_θ corresponds to a $g_\theta \in G_N$ and $t \leq h_\theta(x, y) \Leftrightarrow t \leq \frac{|g_\theta(x) - y|}{|y|}$. Then the set of k points defined by $w_j = (z_j, |y_j|t_j)$ is shattered by $H^+(G_N, l_{MAE})$ because the h'_θ associated in $H_{N,MAE}$ to the g_θ are such that $\forall j, \mathbb{I}_{t \leq h'_\theta(x,y)}(x_j, y_j, |y_j|t_j) = \theta_j$. Therefore

$$V_{MAPE} := VC_{dim}(H^+(G_N, l_{MAPE})) \leq VC_{dim}(H^+(G_N, l_{MAE})) := V_{MAE}.$$

Using theorem 9.4 from [4], we can bound the L^p covering number with a VC-dim based value. If $V_l = VC_{dim}(H^+(G_N, l)) \geq 2$, $p \geq 1$, and $0 < \epsilon < \frac{B_{N,l}}{4}$, then

$$\mathcal{N}_p(\epsilon, H(G_N, l), D) \leq 3 \left(\frac{2eB_{N,l}^p}{\epsilon^p} \log \frac{3eB_{N,l}^p}{\epsilon^p} \right)^{V_l}. \quad (4)$$

When this bound is plugged into equation (3), it shows the symmetry between B_Y and λ as both appears in the relevant $B_{N,l}$.

4.5 Consistency

Mimicking Theorem 10.1 from [4], we can prove a generic consistency result for MAPE ERM learning. Assume given a series of classes of models, $(G_n)_{n \geq 1}$ such that $\bigcup_{n \geq 1} G_n$ is dense in the set of measurable functions from \mathbb{R}^p to \mathbb{R} according to the $L^1(\mu)$ metric for any probability measure μ . Assume in addition that each G_n leads to a finite VC-dim $V_n = VC_{dim}(H^+(G_n, l_{MAPE}))$ and that each G_n is uniformly bounded by B_{G_n} . Notice that those two conditions are compatible with the density condition only if $\lim_{n \rightarrow \infty} v_n = \infty$ and $\lim_{n \rightarrow \infty} B_{G_n} = \infty$.

Assume finally that (X, Y) is such as $|Y| \geq \lambda$ (almost surely) and that $\lim_{n \rightarrow \infty} \frac{v_n B_{G_n}^2 \log B_{G_n}}{n} = 0$, then $L_{MAPE}(\hat{g}_{l_{MAPE}, n})$ converges almost surely to L_{MAPE}^* , which shows the consistency of the ERM estimator for the MAPE.

The proof is based on the classical technique of exponential bounding. Plug in equation (4) into equation (3) gives a bound on the deviation between the empirical mean and the expectation of

$$K(n, \epsilon) = 24 \left(\frac{16eB_n}{\epsilon} \log \frac{24eB_n}{\epsilon} \right)^{v_n} e^{-\frac{n\epsilon^2}{128B_n^2}},$$

with $B_n = 1 + \frac{B_{G_n}}{\lambda}$. Then it is easy to check that the conditions above guarantee that $\sum_{n \geq 1} K(n, \epsilon) < \infty$ for all $\epsilon > 0$. This is sufficient to show almost sure convergence of $L_{MAPE}(\hat{g}_{l_{MAPE}, n}) - L_{MAPE, G_n}^*$ to 0. The conclusion follows from the density hypothesis.

5 Conclusion

We have shown that learning under the Mean Absolute Percentage Error is feasible both on a practical point of view and on a theoretical one. In application contexts where this error measure is adapted (in general when the target variable is positive by design and remains quite far away from zero, e.g. in price prediction for expensive goods), there is therefore no reason to use the Mean Square Error (or another measure) as a proxy for the MAPE. An open theoretical question is whether the symmetry between the upper bound on $|Y|$ for MSE/MAE and the lower bound on $|Y|$ for the MAPE is strong enough to allow results such as Theorem 10.3 in [4] in which a truncated estimator is used to lift the bounded hypothesis on $|Y|$.

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